# The ${ }^{13} \mathrm{C}$ Nuclear Magnetic Resonance Substituent Chemical Shifts of Monosubstituted Cyclohexanes. Interpretation by Characteristic Vector Analysis 

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#### Abstract

The set of literature ${ }^{13} \mathrm{C}$ n.m.r. substituent chemical shifts (s.c.s.) of 43 monosubstituted cyclohexanes has been used in Characteristic Vector Analysis (CVA). All variability in the spectra generated by substituents and their configuration could be explained by three orthogonal characteristic vectors. Experimental s.c.s. are thus linear combinations of three independent phenomena acting at each carbon atom simultanously with varying degrees of intensity. The largest contribution, described by $V_{1 . n}$ is associated with the polarizing power of the substituent through $\sigma$ bonds. The contribution to s.c.s. by $V_{2 . n}$ is due to a field-inductive effect of the substituent. The contribution associated with $V_{3 . n}$ is significant only for $\mathrm{C}-\gamma$ and $\mathrm{C}-\delta$ and results in an upfield or downfield shift due to the conformation of the substituent.


Correlations between reactivity parameters which apply to chemical equilibria and reaction rates and spectroscopic properties such as ${ }^{13} \mathrm{C}$ n.m.r. substituent chemical shifts (s.c.s.), are well known in the literature. ${ }^{1}$ Equilibrium constants or rates of reactions can be satisfactorily predicted by measuring the chemical shift of the derivative in question, provided that the reactivity parameters $\sigma$, and the chemical shifts, reflect electronic ground-state perturbations at the nucleus.

In order to get a reasonable correlation, the Hammett equation has often been replaced by two-parameter equations. ${ }^{1,2}$ The most frequent models used in n.m.r. spectroscopy are $\sigma_{1}-\sigma_{\mathrm{R}}$ by Taft ${ }^{3}$ and Swain and Luptons $\mathscr{F}-\mathscr{R}^{4}{ }^{4}$ The latter was recently re-examined, ${ }^{5}$ and improved $\mathscr{F}$ and $\mathscr{R}$ scales were quoted. The aim of these models was to achieve a separation of polar field and mesomeric contributions to the total substituent effects.
The significance of these treatments is in dispute because too much stress has been laid on the correlation coefficient as a criterion of goodness of fit, and not enough on the number of independent mechanisms of the transmission of substituent effects.

Recently a new interpretation of n.m.r. parameters in aromatic compounds was introduced, based on principal-component analysis. ${ }^{6}$ It was applied to the analysis of 4 -substituted styrenes, ${ }^{6 a, b}$ 2-substituted indanes, ${ }^{6 c}$ chalcones, ${ }^{6 d}$ phenylthienylacrylonitriles, ${ }^{6 e}$ monosubstituted benzenes, ${ }^{6 f}$ and triazenes, ${ }^{6 g}$ and showed that only one or two principal components are required to describe the s.c.s. variations. These principal components were shown to be related to various reactivity parameters of field-inductive and mesomeric character.

However, the analysis of a class of alicyclic compounds of various origins and structures, but with similar principal components, does not seem to have been attempted. Further development of ${ }^{13} \mathrm{C}$ n.m.r. requires a better theoretical understanding of the shielding mechanism ${ }^{7}$ and a consistent picture of the effects of changes in substituents, and conformation, on chemical shift. ${ }^{8}$ The latter approach has been begun by Dalling and Grant. ${ }^{9}$ Interpretation of the significance of postulated relations and application of derived rules to unknown molecules of similar geometry should be based on rigorous standards of multiple regression analysis. ${ }^{1 a}$

This paper is addressed to the problem of interpretation of substituent and conformation effects on the ${ }^{13} \mathrm{C}$ n.m.r. spectra of monosubstituted cyclohexanes using the concept of Characteristic Vector Analysis (CVA).

Mathematical Foundation of Characteristic Vector Analysis. ${ }^{10}$-A data set of measurements of $n$ individuals with $p$ variables is contained in the data matrix $\mathbf{D}(n \times p)$. It is useful to transform this data set into a new one in which the variables are mutually uncorrelated, and to know the number of linearly independent sources of variation in the system which represent the chosen variables. This method allows the orthogonalization of a set of measurements, and enables us to reach an 'economical' description.

CVA thus provides a method of constructing new variables which are pairwise uncorrelated. Each characteristic vector is a linear combination of the observed variables, and these linear functions are chosen to be orthogonal. Thus, the characteristic vectors are composites of the original variables. The number of characteristic vectors existing in characteristic vector space is equal to, or less than, the number of variables in the data set, $p$. The first characteristic vector is defined as that linear combination of variables which has the maximum variance of all linear functions derivable from the given variables. The second characteristic vector is the linear combination of variables having the maximum variance of all linear functions of the given variables that are orthogonal to the first characteristic vector, and so on. The coefficients of the characteristic vectors are called characteristic vector loadings, contained in matrix $\mathbf{L}$ ( $k \times p$ ), while the magnitudes of characteristic vectors upon each of the individuals are called CV scores, contained in matrix $\mathbf{V}(n \times k)$. $k$ Denotes the number of orthogonal characteristic vectors, and usually $k \ll p$.
The transformation of experimental data $d_{i}$ into $V_{i}$, which have the properties of mutual orthogonality and maximum variance, is achieved by defining $V_{i}$ as the linear combination of $d_{i}$ [equation (1)] or in matrix notation [equation (2)].

$$
\begin{gather*}
V_{1, i}=a_{1,1} d_{1, i}+a_{1,2} d_{2, i}+\cdots \cdots \cdots+a_{p, 1} d_{p, i}  \tag{1}\\
(i=1,2, \cdots \cdots n) \\
V_{1, i}=\mathbf{D} a_{1} \tag{2}
\end{gather*}
$$

$V_{1}$ is a vector of $n$ elements (first row in $\mathbf{V}$ ) and $a_{1}$ is a vector of $p$ elements (first column of eigenvector matrix $\mathbf{A}$ calculated from variance-covariance matrix).
The sum of squares of the elements of $V_{1}$ is given by equation (3) with $a_{1}$ chosen so as to maximize (3), under the constraint (4).

$$
\begin{equation*}
\sum_{i=1}^{n} v_{i}^{2}=v_{i} v_{i}^{\prime}=a_{1} \mathbf{D}^{\prime} \mathbf{D} a_{1} \tag{3}
\end{equation*}
$$

$$
\begin{equation*}
a_{1} a_{1}^{\prime}=1 \tag{4}
\end{equation*}
$$

It is true under these conditions that $a_{1}$ is a unit-length eigenvector corresponding to the largest eigenvalue in $\mathbf{D}^{\prime} \mathbf{D}^{11}$ Vectors $a_{2} \cdots a_{p}$ are defined in a similar way.

The variances of the $V_{i}$ can be found from (5). After

$$
\begin{equation*}
V^{\prime} V=\mathbf{A}^{\prime} \mathbf{D}^{\prime} \mathbf{D A} \tag{5}
\end{equation*}
$$

mathematical transformations, and using diagonal matrix $\mathbf{Q}$ of eigenvalues arranged in descending order of magnitude, we finally obtained equation (6).

$$
\begin{equation*}
V^{\prime} V=\mathbf{I D I}=\mathbf{Q} \tag{6}
\end{equation*}
$$

Equation (6) shows that the variances of $v_{i}$ are numerically equal to the elements $q_{i}$ of diagonal matrix $\mathbf{Q}$ of eigenvalues in $\mathbf{D}^{\prime} \mathbf{D}$. The ratio of $q_{i}$ elements to the trace of matrix $\mathbf{Q}$, i.e. equation (7), describes the percent of total data variability explained by the $i$ th characteristic vector.

$$
\begin{equation*}
\frac{q_{i}}{\operatorname{Tr} \mathbf{Q}} \times 100 \% \tag{7}
\end{equation*}
$$

The elements of the variance-covariance matrix $A$ can be converted into elements of characteristic vector loading matrix L.

The transformation of experimental data $d_{i}$ into $v_{i}$ can be explained in geometrical terms, provided that only three properties are considered. If each $d_{i}$ is treated as a geometrical vector, then the rows in data matrix $\mathbf{D}$ are the co-ordinates of $n$ individuals, which can be visualized in Figure 1. It can be seen that the points tend to fall in a single plane. The position of an individual $n$ needs only two co-ordinates rather than the usual three for its description. It is usual to define the new axes such that the first axis is aligned with the longest data dimension, and the second at right angles to it. Thus the projections of the points onto axes $f_{1}$ and $f_{2}$ are the elements of characteristic vectors $V_{1}$ and $V_{2}$ (CV scores) for these points. In Figure 1, $\cos \alpha$ is the loading of the variable represented by the $b_{1}$ axis onto the characteristic vector represented by orthogonal axis $f_{1}$.

A more detailed account of the geometrical model is given by Morrison ${ }^{11 b}$ and Davies. ${ }^{12}$

Choice of Data (Chemical Shifts Data Matrix).-The ${ }^{13} \mathrm{C}$ n.m.r. spectral data of 43 monosubstituted cyclohexanes were selected from the literature. ${ }^{13-17}$ The spectra were recorded in almost uniform conditions, $\mathrm{CFCl}_{3}$ being the solvent. The temperature was kept low ( 180 K ) to provide chemical shifts for axial and equatorial conformers. Only 10 of the compounds reported in ref. 17 were measured at room temperature, and the conformation ratio was not reported. Among monosubstituted cyclohexanes were 17 pairs of axial-equatorial conformers. Substituents available cover the wide range of chemical properties, as shown in Table 1.

CVA Processing of Chemical Shift Data Matrix.-The initial data matrix of the chemical shifts of 43 monosubstituted cyclohexanes for six carbon atoms was constructed $(43 \times 6)$. Since the $\alpha-C(2)$ and $C(6)$ and $\beta-C(3)$ and $C(5)$ carbon atoms were equivalent, the matrix format was then reduced to four chemical shifts per one compound, i.e. $C(1), C(2), C(3)$, and C(4). CVA calculations were performed, and each element of the data matrix could be represented by equation (8) in which $\varepsilon_{n . p}$

$$
\begin{equation*}
d_{n, p}=d_{n}+V_{1, n} l_{1, p}+V_{2, n} l_{2, p}+V_{3, n} l_{3, p}+\varepsilon_{n, p} \tag{8}
\end{equation*}
$$



Figure 1. Geometrical explanation of CVA original axes $a, b, c$ could be replaced by characteristic vector axes $f_{1}$ and $f_{2}$


Figure 2. $a$, plot of $V_{2}$ verus $V_{1}$ and $b$, plot of $V_{3}$ versus $V_{2}: \times$, axial conformers; , equatorial conformers; $\boldsymbol{\Delta}$, undeclared conformation
denotes the residue unexplained by the CVA model at a particular number of the characteristic vector; other symbols were explained previously. The variances of $\varepsilon_{n, p}$ were calculated for each column and each row of the reconstituted data matrix with $k=1,2$, and 3 characteristic vectors. The magnitude of $w_{e . k} \leqslant$ (S.D. $)^{2}$ indicates a satisfactory fit of the experimental data to the CVA model. The plots of $V_{2, n}$ versus $V_{1 . n}$ and $V_{3, n}$ versus $V_{2, n}$ were printed in addition to matrices of characteristic

Table 1. $V_{1}, V_{2}$, and $V_{3}$ characteristic vector scores of equatorial (1-26) and axial (27-43) monosubstituted cyclohexanes

| Subst. | No. | $\bar{d}_{n}$ | $V_{1, n}$ | $V_{2, n}$ | $V_{3 .-7}^{\text {x }}$ | Ref. | No. | $\bar{d}_{n}$ | $V_{1 . n}$ | $V_{2, n}$ | $V_{3, n}^{\mathrm{x}}$ | Ref. |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| CN | 1 | 26.8225 | 2.2397 | 3.4763 | -0.2453 | 13 | 27 | 25.5525 | 2.1931 | 3.4345 | 1.9910 | 13 |
| $\mathrm{C} \equiv \mathrm{CH}$ | 2 | 28.1075 | 2.1852 | 5.3955 | -0.5043 | 13 | 28 | 26.5975 | 3.2592 | 5.6741 | 2.8209 | 13 |
| Me | 3 | 30.7650 | 4.4712 | 6.8119 | -0.5906 | $16 a$ |  |  |  |  |  |  |
|  | 4 | 30.4500 | 4.3504 | 6.5808 | -0.4495 | $16 b$ |  |  |  |  |  |  |
|  | 5 | 31.3000 | 3.8251 | 6.5872 | -0.3261 | 15 | 29 | 27.5950 | 3.1149 | 7.7740 | 4.0520 | 15 |
|  | 6 | 30.9675 | 4.3783 | 6.4553 | -0.7468 | 14 |  |  |  |  |  |  |
| $\mathrm{C}_{4} \mathrm{H}_{9}$ | 7 | 31.2500 | 8.6411 | 3.9126 | $-0.3463$ | 17 |  |  |  |  |  |  |
| NC | 8 | 34.1475 | 21.9247 | 2.8961 | -0.5491 | 13 | 30 | 31.9100 | 22.6847 | 2.9198 | 2.4494 | 13 |
| NCS | 9 | 35.0200 | 24.8912 | 2.6049 | -1.0319 | 13 | 31 | 32.7575 | 24.6522 | 2.8616 | 1.7975 | 13 |
| $\mathrm{N}_{3}$ | 10 | 35.3525 | 28.8402 | -0.0451 | -1.3226 | 13 | 32 | 33.1750 | 28.4221 | 0.6345 | 2.2928 | 13 |
| $\mathrm{O}_{2} \mathrm{CH}$ | 11 | 38.3600 | 39.5662 | -2.2243 | -1.3181 | 13 | 33 | 36.0675 | 38.3230 | -1.2295 | 1.7774 | 13 |
| $\mathrm{O}_{2} \mathrm{CPh}$ | 12 | 38.6350 | 40.2161 | -2.3022 | -1.3377 | 13 | 34 | 36.1975 | 38.7582 | -1.3132 | 1.3930 | 13 |
| $\mathrm{NO}_{2}$ | 13 | 41.5700 | 50.2982 | -4.7986 | -1.8531 | 14 | 35 | 39.5025 | 48.5306 | -4.8337 | 1.2623 | 14 |
| SH | 14 | 32.0575 | 9.2236 | 8.400 | -2.0473 | 13 | 36 | 28.9200 | 10.2679 | 7.7723 | 3.7250 | 13 |
| I | 15 | 31.9125 | 1.2359 | 10.8907 | -3.2794 | 15 | 37 | 31.3150 | 10.3286 | 7.7305 | 1.5458 | 15 |
| Br | 16 | 35.3750 | 21.8926 | 5.5752 | -1.6557 | 17 |  |  |  |  |  |  |
|  | 17 | 36.7050 | 19.9213 | 5.9482 | -3.0989 | 15 | 38 | 35.0550 | 25.4115 | 4.6720 | 2.1957 | 15 |
| $\mathrm{NH}_{2}$ | 18 | 34.4750 | 19.8617 | 5.3167 | -0.7082 | 17 |  |  |  |  |  |  |
| NHMe | 19 | 35.8250 | 27.0133 | 0.7656 | -0.9283 | 17 |  |  |  |  |  |  |
| Cl | 20 | 37.8675 | 27.2603 | 4.2361 | $-2.5995$ | 15 | 39 | 35.6075 | 30.2001 | 3.7377 | 2.4864 | 15 |
| OH | 21 | 38.8750 | 36.9344 | 1.1545 | -0.7593 | 17 |  |  |  |  |  |  |
|  | 22 | 39.6050 | 36.1001 | 0.8411 | -1.2565 | 14 | 40 | 36.3900 | 34.1961 | 1.7532 | 2.7362 | 14 |
| OCOMe | 23 | 38.4000 | 39.7319 | -2.3241 | -0.6046 | 17 |  |  |  |  |  |  |
| $\mathrm{OSiMe}_{3}$ | 24 | 39.4300 | 37.3512 | 1.6569 | -1.5372 | 13 | 41 | 36.3925 | 35.9431 | 2.4225 | 1.9821 | 13 |
| F | 25 | 43.4350 | 56.2483 | -4.3556 | $-1.8000$ | 13 | 42 | 41.1350 | 55.0299 | $-3.6416$ | 1.3082 | 13 |
|  | 26 | 43.4700 | 56.5710 | -4.3050 | -1.7775 | 15 | 43 | 41.2075 | 55.0731 | -3.8192 | 1.2061 | 15 |

Table 2. Main results of CVA calculation for classes $\mathbf{A}+\mathrm{E}, \mathbf{A}$, and $\mathbf{E}$

| Class | Compd. | $\begin{aligned} & \text { TV } \\ & (\%) \\ & (\%) \end{aligned}$ | $\begin{gathered} \mathrm{TV}_{2}{ }^{\text {a }} \\ (\%) \end{gathered}$ | Residual variance in one- and two-component model ${ }^{\text {a }}$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | C (1) | C(2) | C(3) | C(4) |
| A +E | 1-43 | 97.41 | 99.64 | $\begin{aligned} & 0.582165 \\ & 0.003538 \end{aligned}$ | $\begin{array}{r} 15.710210 \\ 0.003002 \end{array}$ | $\begin{aligned} & 4.593660 \\ & 1.578710 \end{aligned}$ | $\begin{aligned} & 4.028676 \\ & 1.878945 \end{aligned}$ |
| E | $1-26$ | 97.24 | 99.95 | 0.708833 0.000278 | $\begin{array}{r} 17.622338 \\ 0.003762 \end{array}$ | $\begin{aligned} & 1.968183 \\ & 0.240155 \end{aligned}$ | $\begin{aligned} & 4.379429 \\ & 0.203821 \end{aligned}$ |
| A | 27-43 | 98.05 | 99.95 | $\begin{aligned} & 0.578151 \\ & 0.000126 \end{aligned}$ | $\begin{array}{r} 12.658196 \\ 0.052697 \end{array}$ | $\begin{aligned} & 6.959858 \\ & 0.120950 \end{aligned}$ | $\begin{aligned} & 0.377059 \\ & 0.346429 \end{aligned}$ |

${ }^{a}$ The inclusion of $V_{3}$ yields $100 \%$ TV and 0.000000 residual variance.
vector scores $\mathbf{V}$, characteristic vector loadings $\mathbf{L}$, and reconstituted data matrices with $V_{1}, V_{2}$, and $V_{3}$.

## Results and Discussion

The initial data matrix of the ${ }^{13} \mathrm{C}$ chemical shifts of 43 monosubstituted cyclohexanes (equatorial, axial, and mixed conformation) was used in CVA calculations. The results show that three characteristic vectors were significant, explaining 97.41, 2.23 , and $0.36 \%$ of the total data variance, respectively. The unexplained residue variance, $w_{\varepsilon, 2}=0.003538$ and 0.003002 for $\mathrm{C}(1)$ and $\mathrm{C}(2)$ respectively, was much lower than the squared standard deviation of chemical shift measurements (S.D.). ${ }^{2}$ This indicates that only two characteristic vectors are required to describe the chemical shift variance of $\mathrm{C}(1)$ and $\mathrm{C}(2)$ in all 43 compounds. On the other hand, $w_{\varepsilon, 2}>1.55$ for $\mathrm{C}(3)$ and $\mathrm{C}(4)$ is not acceptable and calls for a third characteristic vector.
In Figures 2 a and b we show two planar projections of the three-dimensional space. Visually, (Figure 2b) all compounds seem to fall into two clusters, A-axial with positive $V_{3 . n}$ and E equatorial, including 10 compounds of mixed conformation with negative $V_{3, n}$. In addition, Figure 2a shows the tendency to a linear pattern of negative slope between $V_{2}$ and $V_{1}$. The
simple ordination procedures ${ }^{18}$ applied to points of Figure 2 a allow us to distinguish four classes of substituents. The division was done according to the first atom in the substituent of the cyclohexane ring, i.e. $\mathrm{C}, \mathrm{N}, \mathrm{O}$, or halogen.

However the analysis of data for individual clusters does not provide a significantly better CVA model. The best results were achieved when all compounds were divided into axial cyclohexanes (class A) and equatorial cyclohexanes (class E). The main results of CVA calculations for all monosubstituted cyclohexanes (class A $+E$ ) and for classes $\mathbf{A}$ or $\mathbf{E}$ are reported in Table 2.
$\%$ TV is higher by 0.3 and residual variances are significantly lower, especially for carbon atoms $C(3)$ and $C(4)$ in classes $A$ nd $E$ as compared with one class, $A+E$. However, the third characteristic vector must be included into the model to lower the residual variance. CV scores for a particular substituent are practically independent of class for $V_{1} ; V_{2}$ are more dependent on class and $V_{2}$ are very strongly dependent. In Table 1, the values of $V_{1}, V_{2}$, and $V_{3}$ are reported for class A and E compounds.

It can be seen from Figure 2 and the data in Table 1 that values of $V_{1 . n}$ are distributed over a range of 60 units, and values of $V_{2, n}$ over a range of 17 units. In addition the tendency to linear dependence of negative slope can be observed. The

Table 3. Data of characteristic vector loadings for axial (A) and equatorial ( $\mathbf{E}$ ) cyclohexanes

|  |  | C(1) | C(2) | C(3) | C(4) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathrm{L}_{1}$ | E | 0.8486103 | $-0.1200753$ | $-0.3697861$ | -0.3587489 |
|  | A | 0.8490772 | -0.1417468 | $-0.4200637$ | -0.2872668 |
| $L_{2}$ | E | -0.171024 1 | 0.8528165 | $-0.2666185$ | -0.415 1735 |
|  | A | -0.169778 3 | 0.7928476 | -0.5839868 | -0.039 0823 |
| $L_{3}$ | E | -0.024 7257 | 0.0510272 | -0.736 3238 | 0.6700223 |
|  | A | -0.015 5917 | -0.3182775 | -0.482 1886 | 0.8160378 |

magnitude of $V_{1, n}$ increases when the electronegativity of the substituents increases, as a general rule. For example $\mathrm{I}^{-}$of small electronegativity has a very small $V_{1}$ value, and $\mathrm{F}^{-}$of high electronegativity indicates the highest $V_{1}$ value. On the other hand, the magnitude of $V_{2, n}$ decreases as the electronegativity of the substituent increases. Closer examination of the relation between $V_{1, n}$ or $V_{2, n}$ and electronegativity ${ }^{19}$ indicates the random scattering of points. Only a few substituents, including halogens, amino groups, and hydroxy groups, form good straight lines. Alkyls, $\mathrm{C} \equiv \mathrm{CH}, \mathrm{NHMe}, \mathrm{OAc}, \mathrm{NO}_{2}$, and NCS produce very large positive or negative deviations and for the other 5 substituents no data on electronegativity were available. The other substituent parameters reflecting inductive properties (e.g. $\sigma_{1}, \mathscr{F}$ ) gave even worse results.

Hence characteristic vectors are orthogonal, they represent unrelated phenomena. The explanation of their chemical background must be based on observed relations to electronegativity and the magnitude of characteristic vector loading presented in Table 3, and could be explained by existing theory.

Characteristic vector loadings for $V_{1}$ are positive and very large for $\mathrm{C}(1)$ but small and negative for all other carbon atoms. We believe that $V_{1, n}$ describes the polarization power of substituents through the $\sigma$ bond in axial and equatorial cyclohexanes (and perhaps in other alicyclic systems) as visualized by ${ }^{13} \mathrm{C}$ n.m.r. Electronegative substituent-attracting electrons from carbon induce a partial positive charge at $C(1)$ and then induce partial negative charges at $C(2), C(3)$, and $C(4)$ by $\sigma$-inductive nonalternant effect. ${ }^{20,21}$ Thus CV loadings represent the magnitude and sign of those partial charges.

CV loadings for $V_{2}$ are positive and very large for $\mathrm{C}(2)$ but negative for all other carbons, $C$ (3) being more negative. These facts correlate with the relationship between the magnitude of $V_{2}$ and electronegativity, showing that the high sensitivity associated with the first phenomenon is accompanied by weak activity of the second phenomenon, and vice versa. In our opinion, $V_{2, n}$ describes the inductive power of the substituent by field effect ${ }^{2,2}$ as visualized by ${ }^{13} \mathrm{C}$ n.m.r. Field effects induce a positive charge at $C(2)$ through space, and perturb $\sigma$-orbitals by a $\sigma$-inductive nonalternatant effect, inducing negative charges at $\mathrm{C}(3)$ and $\mathrm{C}(4)$. Field effects depend on the dipole moment of the substituent ${ }^{22}$ and could be approximated by $\sigma_{1},{ }^{3}{ }^{\mathscr{F}},{ }^{4,5-23}$ or other similar scale in many applications. However the field effect upon nuclear magnetic shielding depends upon the dipole moment, but with different geometric terms, ${ }^{24}$ and the relationship between the above scales and $V_{2, n}$ is random. We believe that nonproportionality between $V_{1}$ and $V_{2}$ is caused by the nonproportional nature of field and inductive effects, ${ }^{24 a, d}$ especially for polyatomic substituents.

It is difficult to explain the nature of the third mechanism which contributes to the overall shift effect by moderate or weak shifts of the $\gamma$ and $\delta$ carbons, depending on substituent and conformation.

Schneider and Hoppen ${ }^{13}$ reported that electronegative substituents shield C- $\delta$ axial groups in all cases to a significantly
lesser degree (1--2 p.p.m.). CNDO- and INDO-calculated charge densities fail to reproduce this shift. By contrast, frontier orbital treatment ${ }^{25}$ qualitatively predicts an accumulation of electron density at $\mathrm{C}-\delta$ by electronegative substituent. Recent force-field calculations ${ }^{26}$ indicate no significant geometry changes upon equatorial halogen, alkoxy, and alkyl substituents in cyclohexane, but axial substituents cause a change in torsional angles. ${ }^{27}$

A specific role for axial substituents has been proposed in the investigation of $\mathrm{C}-\gamma$ shifts in $\gamma$-gauche conformations of hydrocarbons by sterically induced charge polarization of the $\mathrm{C}-\mathrm{H}$ bond. ${ }^{28}$ The coupled bond angle-torsional angle gauche effect theory ${ }^{29}$ seems to be applicable in the compounds studied. Assuming that CV loadings $L_{3}$ will characterize the sign and relative magnitude of charge in the molecule we come to agreement with the sterically induced charge polarization concept.

We can see from the data in Table 3 the accumulation of negative CV loadings at $C(3)$ and the accumulation of negative loadings weakened by positive CV loadings induced by the third mechanism at $\mathrm{C}(4)$.

Of key importance for a contribution by a third mechanism, however, is the sign and magnitude of the third characteristic vector. $V_{3}$ for axial conformers is always negative for equatorial conformers. $V_{3}$ determines the upfield or downfield shift at $\mathrm{C}(3)$ and $C(4)$ associated with the third mechanism.

Since the sign of $V_{3}$ depends on the orientation of the substituent, we can imagine and examine in future a possible link between the third mechanism and the field gradient shift. ${ }^{24 b}$

We have to keep in mind that the concept of substituent chemical shift presented now differs basically from the previous one. ${ }^{13}$ Now we treat s.c.s. as a linear combination of three independent contributions at each carbon atom described by the general equation (8). The parameters of this equation for various substituents are presented in Table 1, and for carbon atoms in Table 3. The general applicability of these parameters to other alicyclic systems will be discussed in another paper.

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